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Association between process measures and mortality in individuals with opioid use disorders

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ABSTRACT

Background: Individuals with opioid use disorders have high rates of mortality relative to the general population. The relationship between treatment process and mortality is unknown.

Aim: To examine the association between 7 process measures and 12- and 24-month mortality.

Methods: Retrospective cohort study of patients with opioid use disorders who received care from the Veterans Administration between October 2006 and September 2007. Logistic regression models were used to examine the association between 12 and 24-month mortality and 7 patient-level process measures, while risk-adjusting for patient characteristics. Process measures included quarterly physician visits, any opioid use disorder pharmacotherapy, continuous pharmacotherapy, psychosocial treatment, Hepatitis B/C and HIV screening, and no prescriptions for benzodiazepines or opioids. We conducted sensitivity analyses to examine the robustness of our findings to an unobserved confounder.

Results: Among individuals with opioid use disorders, not being prescribed opioids or benzodiazepines, receipt of any psychosocial treatment and quarterly physician visits were significantly associated with lower mortality at both 12 and 24 months, but Hepatitis and HIV screening, and measures related to opioid use disorder pharmacotherapy were not. Sensitivity analyses indicated that the difference in the prevalence of an unobserved confounder would have to be unrealistically large given the observed data, or there would need to be a large effect of the confounder, to render these findings non-significant.

Conclusions and relevance: This is the first study to show an association between process measures and mortality in patients with opioid use disorders and provides initial evidence for their use as quality measures.

1. Introduction

Opioid use disorders (OUDs), including both heroin and prescription opioid use disorders, are associated with high rates of mortality both in the United States and worldwide (Degenhardt et al., 2011; Degenhardt et al., 2014a; Gomes et al., 2014; Martins et al., 2015; Rudd et al., 2016b). Mortality risks for individuals with opioid use disorders range from 6 to 20 times higher than the general population, and opioid-related mortality has increased dramatically during the past two decades, particularly with the increasing misuse of prescription opioids (Centers for Disease Control and Prevention (CDC), 2011; Frenk et al., 2015; Rudd et al., 2016a; Rudd et al., 2016b). Younger individuals are disproportionately affected and bear the greatest burden of premature mortality. In addition to being associated with medical illnesses such as endocarditis and hepatitis, OUDs complicate the treatment of other conditions (e.g., acute and chronic pain, conditions requiring surgical intervention) and are associated with higher costs (Birnbaum et al., 2011; Meyer et al., 2014; Ronan and Herzig, 2016).

Reducing the mortality associated with OUDs is an ongoing public health challenge and an important goal for health care systems. While health care systems have little ability to decrease some causes of premature mortality, such as injuries and homicides, they may influence...
mortality through the quality of the care they deliver. Health care providers can decrease the mortality risk of patients with OUDs through recognition and treatment of worsening chronic health problems, preventing iatrogenic causes of mortality or through influencing patients’ drug use and subsequent risk behaviors by providing effective treatment (Davoli et al., 2007; Gaither et al., 2016a, 2016b; Grossbard et al., 2014).

Understanding the relationship between the quality of care provided and mortality can help health care systems reduce the increased mortality experienced by persons with OUDs. Quality of care is commonly examined using either measures of process, which assess the care provided to the patient, or outcomes, which assess the impact of the care on the patient’s health or functioning. While improved patient outcomes are the gold standard for measuring quality, using outcome-based quality measures is challenging. Outcome data can be expensive and difficult to obtain; do not identify which care processes need to be improved; and outcome measures require risk adjustment. In contrast, process-based measures are more easily operationalized from electronic health records and can provide information about where performance is variable and quality improvement efforts should be targeted. Furthermore, administrative-data based process measures can be reported in real-time, allowing health care systems to take timely corrective action.

While the rationale for specific quality measures usually comes from practice guidelines and/or a synthesis of the literature, process measures should have demonstrated reliability and validity before being used as quality measures to improve performance. There are no reliable and valid quality measures for individuals with OUDs, and measures that have been developed and validated for use in substance use disorders more generally have not been specifically tested in opioid dependent populations (Garnick et al., 2002; Garnick et al., 2009; Harris et al., 2010). This is an important limitation, because disease-specific measures, when they draw attention to specific clinical processes, may be more actionable than generic measures for quality improvement efforts. However, unless process measures are associated with clinically meaningful outcomes, using them to monitor and improve performance will not result in the expected improvements.

Given the importance of mortality as a clinical outcome, we examined the association of 7 process-based measures with 12- and 24-month mortality among persons with OUDs. If these measures are associated with lower mortality, it would provide initial evidence that they could be used by health care systems as part of specific strategies to improve the care provided to individuals with OUDs and to decrease mortality. It would also provide initial evidence for the predictive validity of the measures.

2. Methods

2.1. Overview

This study was approved by the Institutional Review Boards of the Central Arkansas Veterans Healthcare Center and the University of Arkansas for Medical Sciences. The boards waived the requirement for informed consent as it was a minimal risk study. Administrative data was obtained from the Veterans Administration (VA) Medical SAS data sets. Mortality through September 30, 2009 was obtained from the VA Vital Status Mini File.

2.2. Study population

We identified all veterans with OUDs using the International Classification of Diseases (ICD)-9 codes 304.0, 304.7, 305.5 found in the VA Medical SAS datasets, which represent the totality of care provided or paid for by VA in federal fiscal year (FY) 2007. Veterans were included in the study population if within FY2007 their utilization records contained at least one diagnosis code for an OUD, and if they had at least one inpatient episode or two outpatient encounters, one of which was related to OUD, to show active engagement with VA care. Patients were not required to be receiving treatment for their OUD.

2.3. Process measures

Process measures were derived from a comprehensive literature review as part of a larger study. Discrete treatment recommendations were identified and potential measures were reviewed by a panel of technical experts and iteratively revised until a final set of measures of acceptable face validity and feasibility was produced with all necessary technical specifications (Watkins et al., 2011a, 2016, 2011b). The measure related to benzodiazepine and opioid use, and 3-month opioid pharmacotherapy were derived from those previously used to assess the treatment of Massachusetts and Pennsylvania Medicaid enrollees with OUDs (Baxter et al., 2015; Gordon et al., 2015; Stein et al., 2012). Because of the relatively low prevalence of the mortality outcome, we only included the 7 measures that were applicable to a broad set of individuals with OUDs. Receipt of Hepatitis B or C or HIV screening was defined as receiving a blood test for the illness during the observation year; HIV and Hepatitis C screening are recommended by the VA/DoD practice guideline for substance use disorders (U.S. Department of Veterans Affairs, 2016). No opioid and benzodiazepine use, a measure indicating the absence of possible inappropriate prescribing, was defined as not receiving either a prescription for an opioid analgesic or a benzodiazepine during the study period. Any OUD pharmacotherapy was defined as receiving at least one prescription during the study year for an FDA-approved medication for OUD, including oral naltrexone, methadone or buprenorphine. We defined 3-month OUD pharmacotherapy as receiving at least 3 months of OUD pharmacotherapy; for both medication measures we required the methadone to be dispensed through an opioid treatment program to exclude use for pain. Given the importance of psychosocial interventions (U.S. Department of Veterans Affairs, 2016), receipt of any psychosocial treatment was defined as receiving at least one diagnosis-related psychosocial treatment visit in the observation year, including individual and group psychotherapy, family interventions, supported employment, skills training and intensive case management, but excluding opioid substitution therapy. The final measure describes an aspect of continuity of care, continuous care over time (Wierdema et al., 2009), which we defined as receiving at least one physician visit for any diagnosis each quarter over a one-year period. We tested alternative specifications, including restricting the visit to an OUD or SUD diagnosis-related visit. Because the relationships observed were similar, we present data from the least restrictive version of the measure. Measures were assessed during FY 2007.

2.4. Outcome measure

Mortality was obtained from the VA Vital Status Mini File for the period from October 1, 2006 through September 30, 2009. The Mini File contains information from Medicare Vital Status file, the Social Security Administration Death Master file, Patient Treatment File, Veterans Beneficiary Identification Records Locator Subsystem Death File and Fee Basis. Death dates compiled from these combined data sources demonstrated very high sensitivity and exact agreement with dates from the National Death Index (Sohn et al., 2006).

2.5. Covariates

To risk-adjust mortality rates, we used variables available in the administrative data, including age, gender, racial/ethnic background, marital status, rural/urban location (Morrill et al., 1999), and whether the veteran had a service-connected disability for a mental or substance use disorder (SUD). Service-connection status was included because it is associated with illness severity and veterans with a service-connected
disability are given priority access to services. Given that patients with multiple comorbidities show increased healthcare utilization but worse outcomes, a comorbidity measure based on the Charlson-Deyo comorbidity index (Deyo et al., 1992; Klabunde et al., 2000) was used to adjust for mortality risk due to physical health conditions. The index was modified by the VA Information Resource Center for use with mixed inpatient and outpatient data (VA Information Resource Center, 2014). We include for descriptive purposes the number and proportion of the population who had a new treatment episode, defined as having at least one substance use or mental health visit after a period of 5 months with no diagnosis-related visits (Harris et al., 2015).

2.6. Statistical analyses

We examined descriptive statistics for 12- and 24-month mortality outcomes, patient risk-adjustment characteristics, and the measures of treatment process. We restricted analyses to the population of patients who were alive at the end of the observation period for each measure in order to unbiasedly estimate mortality following quality measure-specific landmark times (Dafni, 2011). For our primary analyses examining the overall process-outcomes association for each measure and mortality time point, we fit a logistic regression to model the probability of mortality, including the process measure and patient risk-adjustment characteristics as independent variables. Observations with missing covariate data or mortality information (approximately 4.3% of the population) were omitted from the outcomes analyses. We assessed the strength of association between a measure and mortality by examining the odds ratio of mortality for the measure and its 95% confidence interval (CI). We applied the predictive margins approach to the risk-adjusted logistic regression output to estimate the marginal effect of mortality of receipt of care measured by the process measure, holding constant the risk-adjustment patient characteristics (Grubard and Korn, 1999), and computed the marginal percent reduction in mortality associated with receiving a process measure. We also report the avoidable excess mortality number, which refers to the number of deaths that potentially could have been averted had the patient received the process of care specified by the measure. For a specific measure, the avoidable mortality number was calculated as the product of the difference in mortality rates between those who met and did not meet the measure, and the size of the population of patients who did not receive measured care. Standard errors of model coefficients were adjusted for the clustering of observations within one of 139 service areas. Service areas are geographic regions nested within 21 regionally-defined Veterans Integrated Service Networks.

We performed two secondary analyses. Because the overall association between the process measures and mortality might reflect differences between service areas (Finney et al., 2011), we also examined the within-service area associations by fitting logistic regression models similar to those described above but adding fixed-effect terms for service areas instead of cluster-adjusting for service areas. The estimated odds ratio for these analyses compares mortality risk by receipt of the process measure for patients within the same service area.

2.7. Sensitivity analysis

A complication to examining the association between receipt of care and mortality using observational data is that the amount and type of care patients get could differ based on the severity of their illness in a way unexplained by the measured data on patient risk factors (Lin et al., 1998). We applied a sensitivity analysis approach to evaluate how sensitive our results would be to a hypothetical dichotomous unmeasured confounder, U, that was unavailable in the data and had a positive association with mortality. We implement this by assuming the true logistic regression model should contain an additional term, $b \times U$, where $b$ is the regression coefficient for $U$, the value of a hypothetical unobserved confounder for patient $i$. We examine how large an effect $U$ would need to have to invalidate our statistically significant findings. For each process measure, we examine three scenarios under which $U$ is associated with higher mortality:

1. The magnitude of the effect of $U$ is the size of the average process measure effect across all of the analyses ($OR(U) = \exp(b)$)
2. The magnitude of the effect of $U$ is equal to the maximum process measure effect ($OR(U) = 1.43$)
3. The magnitude of the effect of $U$ exceeds the largest observed effect of the process measure and risk-adjustment variables across all of the analyses ($OR(U) = 1.68$).

These values of $OR(U)$ were chosen since effects of these magnitudes were found in our analyses, making them plausible estimates of the potential size of an unobserved confounder’s effect (Griffin et al., 2012).

3. Results

In FY2007, 32,422 patients with OUDs accessed services provided or paid for by the VA. Table 1 shows their demographic and descriptive characteristics; 96% were male and the average age was 52 (SD = 9). Sixty-eight percent had at least one new SUD treatment episode; 23% had co-occurring PTSD. The unadjusted mortality rate was 3.7% at 12 months (1165 individuals) and 7.2% at 24 months (2272).

Adherence to the process measures ranged from a high of 79% (any psychosocial treatment) to a low of 12% (HIV screening). Ten percent of the population received a prescription benzodiazepine during the study period, 41% received a prescription opioid and 45% received either a prescription benzodiazepine or a prescription opioid. Twenty-five percent received any OUD pharmacotherapy, and 19% received

### Table 1

<table>
<thead>
<tr>
<th>Characteristics of Veterans with Opioid Use Disorder Receiving Care from VHA, FY 2007 (N = 32,422).</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Male, No. (%)</strong></td>
</tr>
<tr>
<td><strong>Age, mean (SD)</strong></td>
</tr>
<tr>
<td><strong>Race/Ethnicity</strong></td>
</tr>
<tr>
<td>White, No. (%)</td>
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<tr>
<td>Black, No. (%)</td>
</tr>
<tr>
<td>Hispanic, No. (%)</td>
</tr>
<tr>
<td>Other/Unknown, No. (%)</td>
</tr>
<tr>
<td><strong>Marital Status</strong></td>
</tr>
<tr>
<td>Married, No. (%)</td>
</tr>
<tr>
<td>Not Married, No. (%)</td>
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<tr>
<td><strong>Patient setting</strong></td>
</tr>
<tr>
<td>Rural, No. (%)</td>
</tr>
<tr>
<td>Urban, No. (%)</td>
</tr>
<tr>
<td>Service connected, No. (%)</td>
</tr>
<tr>
<td><strong>Mental Health Disorder</strong></td>
</tr>
<tr>
<td>Schizophrenia</td>
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<tr>
<td>Bipolar I Disorder</td>
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<tr>
<td>PTSD</td>
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<tr>
<td>Major Depression</td>
</tr>
<tr>
<td>Charlson-Deyo Morbidity Index</td>
</tr>
<tr>
<td>With NTE, No. (%)</td>
</tr>
<tr>
<td>With SUD NTE, No. (%)</td>
</tr>
<tr>
<td>With MH NTE, No. (%)</td>
</tr>
<tr>
<td><strong>Mortality</strong></td>
</tr>
<tr>
<td>12-month, No. (%)</td>
</tr>
<tr>
<td>24-month, No. (%)</td>
</tr>
</tbody>
</table>

* Does not equal 100% due to missing data.
** RUCA code missing for 329 patients.
* New Treatment Episode.

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1 For non-dichotomous predictors age and Charlson index, the odds ratios reflect the effect of age/10 and a change of 0.1 points in the Charlson comorbidity index, respectively.
OUD pharmacotherapy for at least three months. Thirty-three percent saw a physician at least once a quarter.

Fig. 1 shows the risk-adjusted odds ratio estimates of 12- and 24-month mortality for the process measures in the primary outcomes analyses, where the odds ratios are represented as squares and their 95% confidence intervals as horizontal segments. Not being prescribed opioids and/or benzodiazepines, psychosocial treatment, and quarterly physician visits were significantly associated with lower odds of mortality at both 12- and 24-months (p < 0.01 for 12-month quarterly physician visits; p < 0.001 for the other measures and time points). The associations of hepatitis screening, HIV screening, receipt of OUD pharmacotherapy, and 3-month OUD pharmacotherapy with 12- and 24-month mortality were not significant.

Table 2 translates the model results shown in Fig. 1 to predicted probabilities of mortality by receipt of each process measure, and shows the avoidable excess mortality for each measure. Receiving the care described by the measure reduced 12-month mortality by up to 29% (no prescription opioids or benzodiazepines).

Fig. 2 summarizes how large an effect an unobserved confounder would need to have to render non-significant the significant multivariate analysis findings (Lindenauer et al., 2014). Statistical significance depends on the prevalence of U for those who receive the process measure (P1: x-axis), the prevalence among those who do not receive the process measure (P0: y-axis), and the odds ratio of U.

<table>
<thead>
<tr>
<th>Process Measure</th>
<th>12-Month Mortality Rate-Received Measured Care (%)</th>
<th>12-Month Mortality Rate-Did not Receive Measured Care (%)</th>
<th>% Change</th>
<th>Avoidable Excess Mortality, No.</th>
</tr>
</thead>
<tbody>
<tr>
<td>12-Month Mortality</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hepatitis screening</td>
<td>3.8</td>
<td>3.7</td>
<td>2</td>
<td>−27</td>
</tr>
<tr>
<td>HIV screening</td>
<td>3.6</td>
<td>3.7</td>
<td>−5</td>
<td>27</td>
</tr>
<tr>
<td>No prescription opioids or benzodiazepines</td>
<td>3.1</td>
<td>4.3</td>
<td>−29</td>
<td>159</td>
</tr>
<tr>
<td>Receipt of any OUD pharmacotherapy</td>
<td>3.5</td>
<td>3.7</td>
<td>−6</td>
<td>46</td>
</tr>
<tr>
<td>Receipt of OUD pharmacotherapy for at least 3 months</td>
<td>3.5</td>
<td>3.7</td>
<td>−7</td>
<td>50</td>
</tr>
<tr>
<td>Psychosocial treatment</td>
<td>3.5</td>
<td>4.5</td>
<td>−24</td>
<td>53</td>
</tr>
<tr>
<td>Quarterly physician visits</td>
<td>3.2</td>
<td>3.9</td>
<td>−18</td>
<td>142</td>
</tr>
<tr>
<td>24-Month Mortality</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hepatitis screening</td>
<td>7.3</td>
<td>7.2</td>
<td>2</td>
<td>−27</td>
</tr>
<tr>
<td>HIV screening</td>
<td>7.1</td>
<td>7.2</td>
<td>−1</td>
<td>27</td>
</tr>
<tr>
<td>No prescription opioids or benzodiazepines</td>
<td>6.2</td>
<td>8.3</td>
<td>−27</td>
<td>279</td>
</tr>
<tr>
<td>Receipt of any OUD pharmacotherapy</td>
<td>7.0</td>
<td>7.3</td>
<td>−4</td>
<td>69</td>
</tr>
<tr>
<td>Receipt of OUD pharmacotherapy for at least 3 months</td>
<td>7.2</td>
<td>7.2</td>
<td>0</td>
<td>0</td>
</tr>
<tr>
<td>Psychosocial treatment</td>
<td>6.9</td>
<td>8.3</td>
<td>−18</td>
<td>75</td>
</tr>
<tr>
<td>Quarterly physician visits</td>
<td>6.5</td>
<td>7.6</td>
<td>−15</td>
<td>223</td>
</tr>
</tbody>
</table>

***p < .001.  **p < .01.
shading indicates stronger effects of U are required to render the finding non-significant (p > 0.05). Specifically, the dark gray/middle gray/light gray shading indicates combinations of P1 and P0 for which OR(U) = 1.68/1.43/1.30 would render the findings non-significant. Non-shaded areas represent combinations of P0 and P1 for which the significance of the findings holds for the three values of OR(U) examined.

To put these hypothetical differences between P1 and P0 into context, we consider the difference in the prevalence of each dichotomous observed confounder (e.g., covariates in our regression models) by receipt of each measure. The largest observed difference between P1 and P0 across the measures in this study was 16.5 percentage points. For five of the six significant results, the difference between P1 and P0 would need to exceed at least 20 percentage points in order to render findings non-significant given the range of OR(U) examined. To note, the percentage point difference between P1 and P0 required to render the association non-significant is larger than any difference in our observed confounders (e.g., covariates in our regression models) by receipt of each measure. The robustness of the findings holds for the three values of OR(U) examined.

Among individuals with OUDs, receiving 3 of the 7 process measures was associated with lower 12- and 24-month mortality and provides initial evidence for their use as quality measures. Not being prescribed either prescription opioids or benzodiazepines, use of psychological treatment and quarterly physician visits were all associated with lower mortality at 12 and 24 months. It is important to note that while the population was at least minimally engaged with VA care, nearly 17% did not receive any VA outpatient SUD treatment during the measurement year, indicating that the population included individuals who were both in and out of formal SUD treatment. These results are consistent with other research that has shown increased mortality with the use of benzodiazepines for individuals who are on chronic opioids, and decreased mortality in individuals with SUDs on chronic opioid therapy who also received either inpatient or outpatient SUD treatment (Gaither et al., 2016a). Laboratory tests for Hepatitis B, C or HIV were not associated with lower mortality.

Surprisingly, we did not find an association with either OUD pharmacotherapy initiation or maintenance and mortality. Pharmacotherapy for OUD is a well-established component of treatment (Ries et al., 2009) and many studies have shown an association between OUD pharmacotherapy and mortality (Cornish et al., 2010; Degenhardt et al., 2009; Pierce et al., 2016). The lack of association in this population may be due to differences between this and other populations studied, the nature of the services they were receiving, or because of the well-documented increased risk of mortality during times of pharmacotherapy treatment initiation and cessation (Buster et al., 2002; Cornish et al., 2010; Cousins et al., 2016; Degenhardt et al., 2009; Kimber et al., 2015; Pierce et al., 2016; Pierce et al., 2015; Tjagvad et al., 2016). This is also an observational study and causal inferences cannot be made. Our findings underscore the importance of examining the relationship between OUD pharmacotherapy and mortality in different populations receiving services in a range of real world settings, and of creating OUD pharmacotherapy measures that take into account the variations in mortality risk associated with

4. Discussion

Among individuals with OUDs, receiving 3 of the 7 process measures was associated with lower 12- and 24-month mortality and provides initial evidence for their use as quality measures. Not being prescribed either prescription opioids or benzodiazepines, use of psychological treatment and quarterly physician visits were all associated with lower mortality at 12 and 24 months. It is important to note that while the population was at least minimally engaged with VA care, nearly 17% did not receive any VA outpatient SUD treatment during the measurement year, indicating that the population included individuals who were both in and out of formal SUD treatment. These results are consistent with other research that has shown increased mortality with the use of benzodiazepines for individuals who are on chronic opioids, and decreased mortality in individuals with SUDs on chronic opioid therapy who also received either inpatient or outpatient SUD treatment (Gaither et al., 2016a). Laboratory tests for Hepatitis B, C or HIV were not associated with lower mortality.

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Pharmacotherapy initiation, maintenance and cessation. It also underscores the importance of establishing both the reliability and predictive validity of process-based measures before they are adopted as quality measures.

Prior studies have suggested that substantial numbers of individuals treated for OUDs may be prescribed opioids and benzodiazepines (Stein et al., 2016). There are a number of mechanisms whereby prescribing either benzodiazepines or opioids to individuals with OUDs may lead to increased mortality (Charlson et al., 2009; Weisberg et al., 2015). Combination use increases the independent and synergistic sedative properties of both medications, and is the likely mechanism for the increased risk of overdose deaths (Jann et al., 2014; Jones et al., 2012; Vogel et al., 2013). Falls, fractures, automobile accidents and pulmonary-related illnesses are also associated with opioid and/or benzodiazepine use. Avoiding unnecessary prescriptions for opioids and/or benzodiazepines among individuals with OUDs is a prime target for intervention efforts, and could involve multiple complementary initiatives, including prescriber education, such as in the recent Surgeon General letter to prescribers (Surgeon General of the United States, 2016), and efforts to alert providers to such prescribing through electronic health records or prescription drug monitoring programs. Because some individuals with OUDs may be receiving either prescription opioids or benzodiazepines appropriately, this measure should receive further study.

OUDs are frequently co-morbid with medical problems (Centers for Disease Control and Prevention, 2012; Degenhardt et al., 2011; Degenhardt et al., 2014b; Ries et al., 2015), putting individuals at risk for polypharmacy (Dublin et al., 2011; Saunders et al., 2010; Woolcott et al., 2009) and poor care for their medical conditions. The few studies examining the link between utilization and mortality suggest that more service use could increase the receipt of preventive health services as well as the early identification and management of new or worsening physical health problems (Druss et al., 2001; Meyer et al., 2014; Paddock et al., in press; Ronan and Herzig, 2016; Watkins et al., 2016). We found that frequent physician monitoring was associated with decreased mortality. Increasing integration of medical and SUD treatment services may facilitate higher rates of physical health monitoring in the future (Alford et al., 2011; Chung et al., 2016; Melek et al., 2014; Pincus et al., 2015).

Our findings of lower mortality rates with psychosocial treatment are consistent with other studies demonstrating that psychosocial treatment may result in decreased drug use and increased abstinence, and that abstinence from drugs is associated with reductions in mortality risk (Langendam et al., 2001; Peles et al., 2010; Scott et al., 2011). Mortality risk from overdose is lower during SUD treatment (Buster et al., 2002; Pierce et al., 2016), and treatment may permit earlier recognition of relapse or the development of medical complications. As the number of opioid-related deaths continues to increase, clinicians and policy makers need to consider how best to ensure that the population of individuals with OUD are able to receive effective psychosocial treatment.

Our study adds to initial work being done on measure development for OUDs (American Society of Addiction Medicine (ASAM), 2016; Harris et al., 2016; U.S. Department of Health and Human Services, 2017U.S. Department of Health and Human Services, n.d.). The American Psychiatric Association has proposed a measure that captures the proportion of individuals with OUDs who were counselled regarding treatment options, and the Washington Circle group has developed and tested for feasibility an opioid pharmacotherapy initiation measure. The National Committee on Quality Assurance has put forward a related measure on HIV testing in injection drug users. While our measures do not speak specifically to these development processes, none of these measures have been tested for validity, and our results highlight the importance of examining the association of proposed measures with a range of outcomes before accepting them as valid measures of quality. While public reporting of performance can enhance transparency and the value of health care, unless quality measures are both reliable and valid, using them to improve performance poses risks for clinicians, patients and consumers (Austin et al., 2016). Unfortunately there is no agreed standard for how reliable and valid a measure should be before it is accepted and used in public reporting, quality improvement efforts or pay for performance programs.

Strengths of our study include the large, population-based database, the importance of the clinical outcome, and the ease with which these measures can be operationalized. The proposed measures can be populated using administrative data available in many settings, making them feasible to implement and report. Because the measures reflect specific clinical processes, they can provide timely and actionable information to health care systems about where quality needs to be improved. The robustness of our main findings is supported by the sensitivity analysis. Either the difference in the prevalence of an unobserved confounder would have to be unrealistically large given the observed covariate data, or a relatively large effect of an unobserved confounder would be required in order to render these findings non-significant.

Our findings must be viewed within the context of study limitations. The VA system is an organizationally and financially integrated system of care, and we do not know if our results will generalize to care outside of the VA system. We used ICD-9 codes to identify individuals with OUD, and cannot assess the validity of how the codes were applied, which may create selection bias. We note however that the population included all veterans identified and recorded as having an OUD, regardless of whether they were receiving treatment. We cannot distinguish between those who had an OUD related to prescription opioid misuse and those who were using heroin or other illicit opioids, or whether the population included individuals in sustained remission whose ICD-9 diagnosis had not been updated. Observational data analyses can identify associations, but not causal mechanisms. Though sensitivity analyses establish the robustness of our associations for a plausible range of unobserved confounding, results could be sensitive to other types of confounding. Study data is from FY 2007, although all process measures reflect recommended clinical practices that have not changed since FY2007. The association between the continuity of care and the psychosocial treatment measure and mortality was independent of the visit type, suggesting that similar relationships with mortality should be observed today, even if the specific treatment processes have changed.

5. Conclusions

At a time when health care systems are increasingly focused on measuring, assessing, and providing incentives to improve quality (Gled et al., 2015; Patel et al., 2015), the lag in the development of behavioral health measures as compared to physical health measures is concerning (Institute of Medicine, 2006). Our findings of an association between three of our seven process measures and mortality provide support for the predictive validity of these measures and for their use as quality measures. It also furthers the development and validation of quality measures for OUDs.

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Contributors

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